

M3S3/S4 STATISTICAL THEORY II

MULTIVARIATE NORMAL DISTRIBUTION: MARGINALS AND CONDITIONALS

Suppose that vector random variable $\underline{X} = (X_1, X_2, \dots, X_k)^\top$ has a multivariate normal distribution with pdf given by

$$f_{\underline{X}}(\underline{x}) = \left(\frac{1}{2\pi}\right)^{k/2} \frac{1}{|\Sigma|^{1/2}} \exp\left\{-\frac{1}{2}\underline{x}^\top \Sigma^{-1} \underline{x}\right\} \quad (1)$$

where Σ is the $k \times k$ variance-covariance matrix (we can consider here the case where the expected value $\underline{\mu}$ is the $k \times 1$ zero vector; results for the general case are easily available by transformation).

Consider partitioning \underline{X} into two components \underline{X}_1 and \underline{X}_2 of dimensions d and $k - d$ respectively, that is,

$$\underline{X} = \begin{bmatrix} \underline{X}_1 \\ \underline{X}_2 \end{bmatrix}.$$

We attempt to deduce

- (a) the marginal distribution of \underline{X}_1 , and
- (b) the conditional distribution of \underline{X}_2 given that $\underline{X}_1 = \underline{x}_1$.

First, write

$$\Sigma = \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix}$$

where Σ_{11} is $d \times d$, Σ_{22} is $(k - d) \times (k - d)$, $\Sigma_{21} = \Sigma_{12}^\top$, and

$$\Sigma^{-1} = V = \begin{bmatrix} V_{11} & V_{12} \\ V_{21} & V_{22} \end{bmatrix}$$

so that $\Sigma V = I_k$ (I_r is the $r \times r$ identity matrix) gives

$$\begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix} \begin{bmatrix} V_{11} & V_{12} \\ V_{21} & V_{22} \end{bmatrix} = \begin{bmatrix} I_d & 0 \\ 0 & I_{k-d} \end{bmatrix}$$

and more specifically the four relations

$$\Sigma_{11}V_{11} + \Sigma_{12}V_{21} = I_d \quad (2)$$

$$\Sigma_{11}V_{12} + \Sigma_{12}V_{22} = 0 \quad (3)$$

$$\Sigma_{21}V_{11} + \Sigma_{22}V_{21} = 0 \quad (4)$$

$$\Sigma_{21}V_{12} + \Sigma_{22}V_{22} = I_{k-d}. \quad (5)$$

From the multivariate normal pdf in equation (1), we can re-express the term in the exponent as

$$\underline{x}^\top \Sigma^{-1} \underline{x} = \underline{x}_1^\top V_{11} \underline{x}_1 + \underline{x}_1^\top V_{12} \underline{x}_2 + \underline{x}_2^\top V_{21} \underline{x}_1 + \underline{x}_2^\top V_{22} \underline{x}_2. \quad (6)$$

In order to compute the marginal and conditional distributions, we must complete the square in \underline{x}_2 in this expression. We can write

$$\underline{x}^\top \Sigma^{-1} \underline{x} = (\underline{x}_2 - \underline{m})^\top M (\underline{x}_2 - \underline{m}) + \underline{c} \quad (7)$$

and by comparing with equation (6) we can deduce that, for quadratic terms in \underline{x}_2 ,

$$\underline{x}_2^\top V_{22} \underline{x}_2 = \underline{x}_2^\top M \underline{x}_2 \quad \therefore \quad M = V_{22} \quad (8)$$

for linear terms

$$\underline{x}_2^\top V_{21} \underline{x}_1 = -\underline{x}_2^\top M \underline{m} \quad \therefore \quad \underline{m} = -V_{22}^{-1} V_{21} \underline{x}_1 \quad (9)$$

and for constant terms

$$\underline{x}_1^\top V_{11} \underline{x}_1 = \underline{c} + \underline{m}^\top M \underline{m} \quad \therefore \quad \underline{c} = \underline{x}_1^\top (V_{11} - V_{21}^\top V_{22}^{-1} V_{21}) \underline{x}_1 \quad (10)$$

thus yielding all the terms required for equation (7), that is

$$\underline{x}^\top \Sigma^{-1} \underline{x} = (\underline{x}_2 + V_{22}^{-1} V_{21} \underline{x}_1)^\top V_{22} (\underline{x}_2 + V_{22}^{-1} V_{21} \underline{x}_1) + \underline{x}_1^\top (V_{11} - V_{21}^\top V_{22}^{-1} V_{21}) \underline{x}_1, \quad (11)$$

which, crucially, is a sum of two terms, where the first can be interpreted as a function of \underline{x}_2 , given \underline{x}_1 , and the second is a function of \underline{x}_1 only.

Hence we have an immediate factorization of the full joint pdf using the chain rule for random variables;

$$f_{\underline{X}}(\underline{x}) = f_{\underline{X}_2 | \underline{X}_1}(\underline{x}_2 | \underline{x}_1) f_{\underline{X}_1}(\underline{x}_1) \quad (12)$$

where

$$f_{\underline{X}_2 | \underline{X}_1}(\underline{x}_2 | \underline{x}_1) \propto \exp \left\{ -\frac{1}{2} (\underline{x}_2 + V_{22}^{-1} V_{21} \underline{x}_1)^\top V_{22} (\underline{x}_2 + V_{22}^{-1} V_{21} \underline{x}_1) \right\} \quad (13)$$

giving that

$$\underline{X}_2 | \underline{X}_1 = \underline{x}_1 \sim N \left(-V_{22}^{-1} V_{21} \underline{x}_1, V_{22}^{-1} \right) \quad (14)$$

and

$$f_{\underline{X}_1}(\underline{x}_1) \propto \exp \left\{ -\frac{1}{2} \underline{x}_1^\top (V_{11} - V_{21}^\top V_{22}^{-1} V_{21}) \underline{x}_1 \right\} \quad (15)$$

giving that

$$\underline{X}_1 \sim N \left(0, (V_{11} - V_{21}^\top V_{22}^{-1} V_{21})^{-1} \right). \quad (16)$$

But, from equation (3), $\Sigma_{12} = -\Sigma_{11} V_{12} V_{22}^{-1}$, and then from equation (2), substituting in Σ_{12} ,

$$\Sigma_{11} V_{11} - \Sigma_{11} V_{12} V_{22}^{-1} V_{21} = I_d \quad \therefore \quad \Sigma_{11} = (V_{11} - V_{12} V_{22}^{-1} V_{21})^{-1} = (V_{11} - V_{21}^\top V_{22}^{-1} V_{21})^{-1}.$$

Hence, by inspection of equation (16), we conclude that

$$\boxed{\underline{X}_1 \sim N(0, \Sigma_{11})}, \quad (17)$$

that is, we can extract the Σ_{11} block of Σ to define the marginal variance-covariance matrix of \underline{X}_1 .

Using similar arguments, we can define the conditional distribution from equation (14) more precisely. First, from equation (3), $V_{12} = -\Sigma_{11}^{-1} \Sigma_{12} V_{22}$, and then from equation (5), substituting in V_{12}

$$-\Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12} V_{22} + \Sigma_{22} V_{22} = I_{k-d} \quad \therefore \quad V_{22}^{-1} = \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12} = \Sigma_{22} - \Sigma_{12}^\top \Sigma_{11}^{-1} \Sigma_{12}.$$

Finally, from equation (3), taking transposes on both sides, we have that $V_{21} \Sigma_{11} + V_{22} \Sigma_{21} = 0$. Then pre-multiplying by V_{22}^{-1} , and post-multiplying by Σ_{11}^{-1} , we have

$$V_{22}^{-1} V_{21} + \Sigma_{21} \Sigma_{11}^{-1} = 0 \quad \therefore \quad V_{22}^{-1} V_{21} = -\Sigma_{21} \Sigma_{11}^{-1},$$

so we have, substituting into equation (14), that

$$\boxed{\underline{X}_2 | \underline{X}_1 = \underline{x}_1 \sim N \left(\Sigma_{21} \Sigma_{11}^{-1} \underline{x}_1, \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12} \right)}. \quad (18)$$

Thus any marginal, and any conditional distribution of a multivariate normal joint distribution is also multivariate normal, as the choices of \underline{X}_1 and \underline{X}_2 are arbitrary.