M2S1: EXERCISE SHEET 5: SOLUTIONS

1. To compute the covariance need first the marginal expectations of X and Y. The key part of the solution is to realize that the support of the joint density is

$$0 < x < 1, 0 < y < 1, 0 < x + y < 1$$

that is, the "lower left corner" triangle of the unit square, bounded by the three lines x = 0, y = 0, x + y = 1. Now, for 0 < x < 1,

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) \, dy = \int_0^{1-x} cxy(1-x-y) \, dy = cx \int_0^{1-x} y(1-x-y) \, dy$$
$$= cx(1-x)^3 \int_0^1 t(1-t) \, dt \qquad (t=y/(1-x))$$
$$= \frac{c}{6}x(1-x)^3 \qquad 0 < x < 1$$

and

$$\int_{-\infty}^{\infty} f_X(x) \ dx = \int_0^1 \frac{c}{6} x (1 - x)^3 \ dx = 1 \Longrightarrow c = 120$$

and hence

$$f_X(x) = 20x(1-x)^3$$
 $0 < x < 1$ $\therefore E_{f_X}[X] = \int_0^1 20x^2(1-x)^3 dx = \frac{1}{3}$

and, by symmetry, $f_Y(y) = 20y(1-y)^3$ (0 < y < 1), $E_{f_Y}[Y] = \frac{1}{3}$ by symmetry of form of the joint pdf. Also

$$\begin{split} \mathbf{E}_{f_{X,Y}}[XY] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_{X,Y}(x,y) \ dx dy = \int_{0}^{1} \left\{ \int_{0}^{1-y} 120x^{2}y^{2}(1-x-y) \ dx \right\} dy \\ &= \int_{0}^{1} 120y^{2} \left\{ \int_{0}^{1-y} x^{2}(1-x-y) \ dx \right\} dy \\ &= \int_{0}^{1} 120y^{2} \left[\frac{x^{3}}{3}(1-y) - \frac{x^{4}}{4} \right]_{0}^{1-y} \ dy \\ &= \int_{0}^{1} 10y^{2}(1-y)^{4} \ dy \\ &= 10 \left[\frac{y^{3}}{3} - y^{4} + \frac{6y^{5}}{5} - \frac{4y^{6}}{6} + \frac{y^{7}}{7} \right]_{0}^{1} = 10 \left(\frac{1}{3} - 1 + \frac{6}{5} - \frac{2}{3} + \frac{1}{7} \right) = \frac{2}{21} \end{split}$$

and hence

$$\operatorname{Cov}_{f_{X,Y}}[X,Y] = \operatorname{E}_{f_{X,Y}}[XY] - \operatorname{E}_{f_X}[X] \cdot \operatorname{E}_{f_Y}[Y] = \frac{2}{21} - \frac{1}{3} \cdot \frac{1}{3} = -\frac{1}{63}$$

2. (a) Put U = X/Y and Z = X; the inverse transformations are therefore X = Z and Y = Z/U, and note that the new variables are constrained by $0 < Z < \min\{U, 1\}$, as Y < 1. In terms of the multivariate transformation theorem, we have transformation functions defined by

$$g_1(t_1, t_2) = t_1/t_2$$
 $g_1^{-1}(t_1, t_2) = t_2$ $g_2(t_1, t_2) = t_1$ $g_2^{-1}(t_1, t_2) = t_2/t_1$

and the Jacobian of the transformation is given by

$$J(u,z) = \begin{vmatrix} 0 & 1 \\ -z/u^2 & 1/u \end{vmatrix} = \frac{z}{u^2}$$

and hence

 $f_{U,Z}(u,z) = f_{X,Y}(z,z/u) \ z/u^2 = z/u^2 \qquad (u,z) \in \mathbb{U}^{(2)} = \{(u,z) : 0 < z < \min\{u,1\}, u > 0\}$ and zero otherwise, and so

$$f_U(u) = \int_{-\infty}^{\infty} f_{U,Z}(u,z) \ dz = \int_{0}^{\min\{u,1\}} z/u^2 \ dz = \frac{(\min\{u,1\})^2}{2u^2} \qquad u > 0.$$

Put $V = -\log(XY)$ and $Z = -\log X$; the inverse transformations are therefore $X = e^{-Z}$ and $Y = e^{-(v-z)}$, and note that 0 < Z < V. In terms of the theorem, we have transformation functions defined by

$$g_1(t_1, t_2) = -\log(t_1 t_2)$$
 $g_1^{-1}(t_1, t_2) = e^{-t_2}$
 $g_2(t_1, t_2) = -\log t_1$ $g_2^{-1}(t_1, t_2) = e^{-(t_1 - t_2)}$

and the Jacobian of the transformation is given by

$$J(v,z) = \begin{vmatrix} 0 & -e^{-z} \\ -e^{-(v-z)} & e^{-(v-z)} \end{vmatrix} = e^{-v}$$

and hence

$$f_{V,Z}(v,z) = f_{X,Y}(e^{-z}, e^{-(v-z)}) e^{-v} = e^{-v}$$
 $(v,z) \in \mathbb{V}^{(2)} = \{(v,z) : 0 < z < v < \infty\}$

and zero otherwise, and so

$$f_V(v) = \int_{-\infty}^{\infty} f_{V,Z}(v,z) \ dz = \int_{0}^{v} e^{-v} \ dz = ve^{-v} \qquad v > 0$$

and zero otherwise. Note that we can attempt the joint transformation by setting

$$\begin{array}{c} U = X/Y \\ V = -\log(XY) \end{array} \iff \begin{array}{c} X = U^{1/2}e^{-V/2} \\ Y = U^{-1/2}e^{-V/2} \end{array}$$

note that, as X and Y lie in (0,1) we have XY < X/Y and XY < Y/X, giving constraints $e^{-V} < U$ and $e^{-V} < 1/U$, so that $0 < e^{-V} < \min\{U, 1/U\}$. The Jacobian of the transformation is

$$J(u,v) = \begin{vmatrix} \frac{u^{-1/2}e^{-v/2}}{2} & -\frac{u^{1/2}e^{-v/2}}{2} \\ -\frac{u^{-3/2}e^{-v/2}}{2} & -\frac{u^{-1/2}e^{-v/2}}{2} \end{vmatrix} = u^{-1}e^{-v}/2.$$

M2S1 EXERCISES 5 SOLUTIONS: page 2 of 6

Hence

$$f_{U,V}(u,v) = u^{-1}e^{-v}/2$$
 $0 < e^{-v} < \min\{u, 1/u\}, u > 0$

The corresponding marginals are given below: let $g(y) = -\log(\min\{u, 1/u\})$, then

$$f_U(u) = \int_{-\infty}^{\infty} f_{U,V}(u,v) \ dv = \int_{g(y)}^{\infty} \frac{e^{-v}}{2u} \ dv = \left[-\frac{e^{-v}}{2u} \right]_{g(y)}^{\infty} = \frac{\min\{u, 1/u\}}{2u} \quad u > 0$$

$$f_V(v) = \int_{-\infty}^{\infty} f_{U,V}(u,v) \ du = \int_{e^{-v}}^{e^v} \frac{e^{-v}}{2u} \ du = \left[\frac{\log u}{2} e^{-v}\right]_{e^{-v}}^{e^v} = ve^{-v} \qquad v > 0$$

(b) Now let

$$V = X + Y$$
 \iff $X = \frac{V + Z}{2}$ $Z = X - Y$ \Rightarrow $Y = \frac{V - Z}{2}$

and the Jacobian of the transformation is 1/2. The transformed variables take values on the square A in the (V, Z) plane with corners at (0,0), (1,1), (2,0) and (1,-1) bounded by the lines z=-v, z=v and z=v-2. Then

$$f_{V,Z}(v,z) = \frac{1}{2} \qquad (v,z) \in A$$

and zero otherwise (hint: sketch the square A). Hence, integrating in horizontal strips in the (V, Z) plane,

$$f_Z(z) = \int_{-\infty}^{\infty} f_{V,Z}(v,z) \ dv = \begin{cases} \int_{-z}^{2+z} \frac{1}{2} \ dv = 1+z & -1 < z \le 0 \\ \int_{-z}^{2-z} \frac{1}{2} \ dv = 1-z & 0 < z < 1 \end{cases}$$

3. The transformations are

$$Y_1 = \frac{X_1}{X_1 + X_2 + X_3}$$
 $X_1 = Y_1 Y_3$ $Y_2 = \frac{X_1}{X_1 + X_2 + X_3}$ \iff $X_2 = Y_2 Y_3$ $X_3 = Y_3 (1 - Y_1 - Y_2)$

which gives Jacobian

$$J(y_1, y_2, y_3) = \begin{vmatrix} y_3 & 0 & y_1 \\ 0 & y_3 & y_2 \\ -y_3 & -y_3 & (1 - y_1 - y_2) \end{vmatrix} = y_3^2$$

Hence the joint pdf is given by

$$\begin{split} f_{Y_1,Y_2,Y_3}(y_1,y_2,y_3) &= f_{X_1,X_2,X_3}(y_1y_3,y_2y_3,y_3(1-y_1-y_2))J(y_1,y_2,y_3) \\ &= c_1y_1y_3 \exp\left\{-y_1y_3\right\} \ c_2y_2^2y_3^2 \exp\left\{-y_2y_3\right\} \ c_3y_3^3(1-y_1-y_2)^3 \exp\left\{-y_3(1-y_1-y_2)\right\} \ y_3^2 \\ &= c_1c_2c_3y_1y_2^2(1-y_1-y_2)^3 \ y_3^8 \exp\left\{-y_3\right\} = f_{Y_1,Y_2}(y_1,y_2)f_{Y_3}(y_3) \end{split}$$

where

$$f_{Y_1,Y_2}(y_1,y_2) \propto y_1 y_2^2 (1-y_1-y_2)^3$$
 and $f_{Y_3}(y_3) \propto y_3^8 \exp\{-y_3\}$

Hence $Y_3 \sim Gamma(9,1)$; the transformations give the constraints $0 < Y_1, Y_2 < 1$ and $0 < Y_1 + Y_2 < 1$, and $Y_3 > 0$. Now

$$f_{Y_1}(y_1) = \int_{-\infty}^{\infty} f_{Y_1,Y_2}(y_1,y_2) \, dy_2 = \int_{0}^{1-y_1} cy_1 y_2^2 (1-y_1-y_2)^3 \, dy_2 = cy_1 (1-y_1)^6 \int_{0}^{1} t^2 (1-t)^3 \, dt \quad (t = y_2/(1-y_1))^6 \int_{0}^{1} t^2 (1-t)^3 \, dt$$

and hence

$$f_{Y_1}(y_1) \propto y_1(1-y_1)^6 \Longrightarrow Y_1 \sim Beta(2,7), \ f_{Y_1}(y_1) = 336y_1(1-y_1)^6 \qquad 0 < y_1 < 1$$

and hence

$$\mathrm{E}_{f_{Y_1}}[\ Y_1\] = \frac{2}{2+7} = \frac{2}{9}$$

as the expectation of a $Beta(\alpha, \beta)$ distribution is $\alpha/(\alpha + \beta)$ from notes.

4. (a) Put U = X/Y and V = Y; the inverse transformations are therefore X = UV and Y = V. In terms of the multivariate transformation theorem, we have transformation functions defined by

$$g_1(t_1, t_2) = t_1/t_2$$
 $g_1^{-1}(t_1, t_2) = t_1t_2$ $g_2(t_1, t_2) = t_2$ $g_2^{-1}(t_1, t_2) = t_2$

and the Jacobian of the transformation is given by

$$|J(u,v)| = \begin{vmatrix} v & u \\ 0 & 1 \end{vmatrix} = |v|$$

and hence

$$f_{U,V}(u,v) = f_{X,Y}(uv,v) \ |v| = \left(\frac{1}{2\pi}\right) \exp\left\{-\frac{1}{2}(u^2v^2 + v^2)\right\} |v| \qquad (u,v) \in \mathbb{R}^2$$

and zero otherwise, and so, for any real u,

$$f_U(u) = \int_{-\infty}^{\infty} f_{U,V}(u,v) \ dv = \int_{-\infty}^{\infty} \left(\frac{1}{2\pi}\right) \exp\left\{-\frac{1}{2}(u^2v^2 + v^2)\right\} |v| \ dv$$

$$= \left(\frac{1}{\pi}\right) \int_{0}^{\infty} v \exp\left\{-\frac{v^2}{2}(1+u^2)\right\} \ dv \qquad \text{integrand is even func}$$

$$= \left(\frac{1}{\pi}\right) \left[-\frac{1}{(1+u^2)} \exp\left\{-\frac{v^2}{2}(1+u^2)\right\}\right]_{0}^{\infty} = \frac{1}{\pi(1+u^2)} \qquad \text{by direct integration}$$

(b) Now put $T = X/\sqrt{S/\nu}$ and R = S; the inverse transformations are therefore $X = T\sqrt{R/\nu}$ and S = R. In terms of the multivariate transformation theorem, we have transformation functions from $(X, S) \to (T, R)$ defined by

$$g_1(t_1, t_2) = t_1/\sqrt{t_2/\nu}$$
 $g_1^{-1}(t_1, t_2) = t_1\sqrt{t_2/\nu}$ $g_2(t_1, t_2) = t_2$ $g_2^{-1}(t_1, t_2) = t_2$

and the Jacobian of the transformation is given by

$$|J(t,r)| = \begin{vmatrix} \sqrt{\frac{r}{\nu}} & \frac{t}{2\sqrt{r\nu}} \\ 0 & 1 \end{vmatrix} = \left| \sqrt{\frac{r}{\nu}} \right| = \sqrt{\frac{r}{\nu}}$$

and hence

$$f_{T,R}(t,r) = f_{X,S}\left(t\sqrt{\frac{r}{\nu}},r\right)\sqrt{\frac{r}{\nu}} = f_X\left(t\sqrt{\frac{r}{\nu}}\right) f_S\left(r\right)\sqrt{\frac{r}{\nu}} \qquad t \in \mathbb{R}, s \in \mathbb{R}^+$$

and zero otherwise, and so, for any real t,

$$f_{T}(t) = \int_{-\infty}^{\infty} f_{T,R}(t,r) dr$$

$$= \int_{0}^{\infty} \left(\frac{1}{2\pi}\right)^{1/2} \exp\left\{-\frac{rt^{2}}{2\nu}\right\} c(\nu)r^{\nu/2-1}e^{-r/2}\sqrt{\frac{r}{\nu}} dr$$

$$= \left(\frac{1}{2\pi}\right)^{1/2} \frac{c(\nu)}{\sqrt{\nu}} \int_{0}^{\infty} r^{(\nu+1)/2-1} \exp\left\{-\frac{r}{2}\left(1+\frac{t^{2}}{\nu}\right)\right\} dr$$

$$= \left(\frac{1}{2\pi}\right)^{1/2} \frac{c(\nu)}{\sqrt{\nu}} \left(1+\frac{t^{2}}{\nu}\right)^{-(\nu+1)/2} \int_{0}^{\infty} z^{(\nu+1)/2-1} \exp\left\{-\frac{z}{2}\right\} dz \text{ setting } z = r\left(1+\frac{t^{2}}{\nu}\right)$$

$$= \left(\frac{1}{2\pi}\right)^{1/2} \frac{c(\nu)}{\sqrt{\nu}} \left(1+\frac{t^{2}}{\nu}\right)^{-(\nu+1)/2} \frac{1}{c(\nu+1)} \text{ integrand is a pdf}$$

We also see/deduce that f_S is a $Gamma(\nu/2, 1/2)$ or $Chiquared(\nu)$ density, and that the normalizing constant $c(\nu)$ is given by

$$c(\nu) = \frac{\left(\frac{1}{2}\right)^{\nu/2}}{\Gamma\left(\frac{\nu}{2}\right)} \implies f_T(t) = \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\Gamma\left(\frac{\nu}{2}\right)} \left(\frac{1}{\pi\nu}\right)^{1/2} \frac{1}{(1+t^2/\nu)^{(\nu+1)/2}}$$

which is the $Student(\nu)$ density.

5. We have

$$f_{X|Y}(x|y) = \sqrt{\frac{y}{2\pi}} \exp\left\{-\frac{yx^2}{2}\right\} \qquad x \in \mathbb{R} \qquad f_Y(y) = c(\nu)y^{\nu/2 - 1}e^{-\nu y/2} \qquad y \in \mathbb{R}^+$$

where ν is a positive integer, so that $X|Y=y\sim N(0,y^{-1})$ and $Y\sim Gamma(\nu/2,\nu/2)$, and the normalizing constant $c(\nu)$ is given by

$$c(\nu) = \frac{\left(\frac{\nu}{2}\right)^{\nu/2}}{\Gamma\left(\frac{\nu}{2}\right)}$$

Now, by the chain rule

$$f_{X,Y}(x,y) = f_{X|Y}(x|y)f_Y(y)$$
 $x \in \mathbb{R}, y \in \mathbb{R}^+$

and zero otherwise, and so, for any real x,

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) \, dy$$

$$= \int_{0}^{\infty} \sqrt{\frac{y}{2\pi}} \exp\left\{-\frac{yx^2}{2}\right\} \frac{\left(\frac{\nu}{2}\right)^{\nu/2}}{\Gamma\left(\frac{\nu}{2}\right)} y^{\nu/2-1} e^{-\nu y/2} \, dy$$

$$= \frac{1}{\sqrt{2\pi}} \frac{\left(\frac{\nu}{2}\right)^{\nu/2}}{\Gamma\left(\frac{\nu}{2}\right)} \int_{0}^{\infty} y^{(\nu+1)/2-1} \exp\left\{-\frac{y}{2}\left(\nu+x^2\right)\right\} \, dy$$

$$= \frac{1}{\sqrt{2\pi}} \frac{\left(\frac{\nu}{2}\right)^{\nu/2}}{\Gamma\left(\frac{\nu}{2}\right)} \frac{\Gamma\left(\frac{\nu+1}{2}\right)}{\left(\frac{1}{2}\left(\nu+x^2\right)\right)^{(\nu+1)/2}} \quad \text{integrand } \propto \text{a Gamma pdf}$$

Therefore f_X is given by

$$f_X(x) = \frac{\Gamma(\frac{\nu+1}{2})}{\Gamma(\frac{\nu}{2})} \left(\frac{1}{\pi\nu}\right)^{1/2} \frac{1}{(1+x^2/\nu)^{(\nu+1)/2}}$$

which is again the $Student(\nu)$ density.

Exercises 5 and 6 give the two alternative ways of specifying the Student-t distribution, either as a function of independent Normal and Gamma/Chi-squared variables, or as the marginal obtained by "scale-mixing" a Normal distribution by a Gamma distribution (that is, rather than having a fixed variance $\sigma^2 = 1/Y$, we regard Y as a random variable having a Gamma distribution, so that (X,Y) have a joint distribution

$$f_{X,Y}(x,y) = f_{X|Y}(x|y)f_Y(y)$$

from which we calculate $f_X(x)$ by integration.