

M1S : EXERCISE SHEET 7 : SOLUTIONS

1. (i) Density function must integrate to 1 over $\mathbb{X} = [0, 1]$, so

$$\int_0^1 f_X(x) dx = 1 \implies \int_0^1 cx^2 + x dx = 1 \implies \left[c\frac{x^3}{3} + \frac{x^2}{2} \right]_0^1 = 1 \implies c = \frac{3}{2}$$

(ii) Distribution function F_X given for $0 \leq x \leq 1$ by

$$F_X(x) = \int_0^x f_X(t) dt = \frac{x^3 + x^2}{2}$$

and $F_X(x) = 0$ for $x < 0$, and $F_X(x) = 1$ for $x > 1$.

(iii) $P[X < 1/2] = F_X(1/2) = 3/16$.

(iv) From definition of conditional probability

$$P[X > 1/2 \mid X > 1/4] = \frac{P[X > 1/2, X > 1/4]}{P[X > 1/4]} = \frac{P[X > 1/2]}{P[X > 1/4]} = \frac{1 - F_X(1/2)}{1 - F_X(1/4)} = \frac{104}{123}$$

(v) $Y \sim \text{Binomial}(200, \theta)$ where

$$\theta = P[Y_i = 1] = P[X < 1/6] = F_X(1/6) = \frac{7}{432}$$

Then

$$P[Y \leq 3] = P[Y = 0] + P[Y = 1] + P[Y = 2] + P[Y = 3] = 0.593$$

2. Density function must integrate to 1 over $\mathbb{X} = [-1, 1]$, so

$$\int_{-1}^1 f_X(x) dx = 1 \implies c = \frac{3}{4}$$

Distribution function F_X given for $-1 \leq x \leq 1$ by

$$F_X(x) = \int_{-1}^x f_X(t) dt = \left[\frac{3x}{4} - \frac{x^3}{4} \right]_{-1}^1 = \frac{3}{4} \left(x - \frac{x^3}{3} \right) + \frac{2}{3}$$

and $F_X(x) = 0$ for $x < -1$, and $F_X(x) = 1$ for $x > 1$.

3.(i) Density function must integrate to 1 over $\mathbb{X} = [0, \pi/2]$, so

$$\int_0^{\pi/2} f_X(x) dx = 1 \implies \int_0^{\pi/2} cx(\pi - x) dx = 1 \implies c = \frac{12}{\pi^3}$$

(ii) $\mathbb{X} = [0, \pi/2] \implies \mathbb{Y} = [0, 1/2]$, and distribution function F_Y given by

$$F_Y(y) = P[Y \leq y] = P[\sin X \leq 2y] = P[X \leq \sin^{-1}(2y)] = F_X(\sin^{-1}(2y))$$

as \sin is monotone on $[0, \pi/2]$. Hence by differentiation with respect to y ,

$$f_Y(y) = \frac{2}{\sqrt{1 - 4y^2}} f_X(\sin^{-1}(2y)) \quad 0 \leq y \leq 1/2$$

and zero otherwise.

4. By differentiation, $f_X(x) = 2xe^{-x^2}$, $x > 0$, and zero otherwise

5. Need to consider ranges of integration carefully;

$$F_X(x) = \begin{cases} \int_0^x t \, dt & = x^2/2 & 0 \leq x \leq 1 \\ \int_0^1 t \, dt + \int_1^x (2-t) \, dt & = 2x - x^2/2 - 1 & 1 \leq x \leq 2 \end{cases}$$

and $F_X(x) = 0$ for $x < 0$, and $F_X(x) = 1$ for $x > 2$. Hence $P[0.8 < X < 1.2] = F_X(1.2) - F_X(0.8) = 0.36$.

6.(i) $P[T > 3 | A] = 1 - P[T \leq 3 | A] = 1 - \Phi((3-2)/(3/4)) = 1 - \Phi(4/3)$

(ii) By the Theorem of Total Probability

$$P[T > 3] = P[T > 3 | A]P(A) + P[T > 3 | B]P(B) = (1 - \Phi(4/3)) \times \frac{3}{10} + (1 - \Phi(-4/3)) \times \frac{7}{10}$$

(iii) By Bayes Theorem

$$P[A | T > 3] = \frac{P[T > 3 | A]P(A)}{P[T > 3]}$$

so obtain final answer by evaluating $\Phi(4/3)$ via statistical tables, and noting that $\Phi(-4/3) = 1 - \Phi(4/3)$.

7. (i) $X \sim \text{Exponential}(1) \implies F_X(x) = 1 - e^{-x}$ for $x > 0$, so $F_X(x) = 1/2 \implies x = \log 2$.

(ii) $\log X \sim \text{Normal}(\mu, \sigma^2)$, so

$$\begin{aligned} F_X(x) = 1/2 &\implies P[X \leq x] = 1/2 \implies P[(\log X - \mu)/\sigma \leq (\log x - \mu)/\sigma] = 1/2 \\ &\implies \Phi((\log x - \mu)/\sigma) = 1/2 \implies (\log x - \mu)/\sigma = 0 \implies \log x = \mu \implies x = e^\mu \end{aligned}$$

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By differentiation

$$\begin{aligned} \text{(a)} \quad & \frac{e^x}{(1+e^x)^2} & x \in \mathbb{R} \\ \text{(b)} \quad & \alpha \lambda x^{\alpha-1} \exp\{-\lambda x^\alpha - e^{-\lambda x^\alpha}\} & x \in \mathbb{R}^+ \\ \text{(c)} \quad & \alpha \lambda x^{\alpha-1} e^{-\lambda x^\alpha} & x \in \mathbb{R}^+ \end{aligned}$$

For part (ii), by careful integration

$$F_X(x) = \begin{cases} 0 & x < 0 \\ x^2 & 0 \leq x < 1/2 \\ 6x - 3x^2 - 2 & 1/2 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$$