

M1S : EXERCISE SHEET 2 : SOLUTIONS

1.(a) Partition according to disease status : $(S \cap T') \cup (S' \cap T)$

(b) Partition first into correctly classified sufferers/non-sufferers

$$\begin{aligned} \text{(i) Sufferers} & : (D \cap (X \cup T) \cap S) \\ \text{(i) Non-sufferers} & : (((D \cap (X \cup T)') \cup D') \cap S') \end{aligned}$$

(i) follows as sufferers must have the correct doctor's diagnosis, and at least one the tests indicating the disease.

(ii) follows as non-sufferers can either receive the incorrect doctor's diagnosis (and be declared as sufferers), but be classified correctly as they receive (correct) negative X-ray and test results, or merely receive a (correct) negative test diagnosis from the doctor.

2. Let G ="exactly one occurs". Then $G = (E \cap F') \cup (E' \cap F)$, and by Axiom (III) $P(G) = P(E \cap F') + P(E' \cap F)$. Also,

$$\begin{aligned} E &= (E \cap F) \cup (E \cap F') \implies P(E) = P(E \cap F) + P(E \cap F') \\ F &= (E \cap F) \cup (E' \cap F) \implies P(F) = P(E \cap F) + P(E' \cap F) \end{aligned}$$

so therefore

$$P(G) = P(E) + P(F) - 2P(E \cap F).$$

3. (a) FALSE : (e.g. $E_1 = \emptyset$, $E_2 = \Omega$).

(b) FALSE : (e.g. coin toss, let $E_1 = E_2 = \{H\}$).

(c) TRUE : $E_1 \subseteq E_2 \implies P(E_1) \leq P(E_2) = 1 - P(E_2') = 1 - P(E_1) \implies 2P(E_1) \leq 1$.

(d) TRUE : $P(E_1 \cup E_2) = P(E_1) + P(E_2) - P(E_1 \cap E_2) \geq (1 - x_1) + (1 - x_2) - 1 = 1 - x_1 - x_2$.

4. For general events E and F ,

(a) $F \equiv (E \cap F) \cup (E' \cap F)$, so by Axiom (III)

$$P(F) = P(E \cap F) + P(E' \cap F) \implies P(E' \cap F) = P(F) - P(E \cap F)$$

(b) $E \cup F \equiv E \cup (E' \cap F)$, so by Axiom (III)

$$P(E \cup F) = P(E) + P(E' \cap F) = P(E) + P(F) - P(E \cap F)$$

(c) $E \subseteq F \implies F \equiv E \cup (E' \cap F)$, so by Axiom (III), as $P(E' \cap F) \geq 0$,

$$P(F) = P(E) + P(E' \cap F) \geq P(E)$$

(d) *Bonferroni Inequality*: From (ii), as $P(E \cup F) \leq 1$,

$$P(E \cap F) = P(E) + P(F) - P(E \cup F) \geq P(E) + P(F) - 1$$

5. (a) $E' \cup F' = (E \cap F)' \implies P(E' \cup F') = 1 - P(E \cap F) = 1 - z$.

(b) $F = (E \cap F) \cup (E' \cap F)$, so $P(F) = P(E \cap F) + P(E' \cap F)$, so $P(E' \cap F) = y - z$.

(c) $E' \cup F = E' \cup (E \cap F) \implies P(E' \cup F) = P(E') + P(E \cap F) = 1 - x + z$.

(d) $E' \cap F' = (E \cup F)' \implies P(E' \cap F') = 1 - P(E \cup F) = 1 - x - y + z$.

6. $E \cup E' = \Omega$, so by Axioms (II) and (III), $P(E) + P(E') = P(\Omega) = 1$. Hence $P(E') = 1 - P(E)$. But, by Axiom (I)*, $P(E') \geq 0$, hence $P(E) \leq 1$.

7. (a) True for $n = 2$ as $P(A \cup B) = P(A) + P(B) - P(A \cap B) \leq P(A) + P(B)$. So assume true for $n = k$; then

$$P(E_1 \cup \dots \cup E_k \cup E_{k+1}) \leq P(E_1 \cup \dots \cup E_k) + P(E_{k+1}) = \sum_{i=1}^{k+1} P(E_i)$$

and hence true for $n = k + 1$.

(b) As for (a) except

$$1P(E_1 \cup \dots \cup E_k \cup E_{k+1}) = P(E_1 \cup \dots \cup E_k) + P(E_{k+1}) - P((E_1 \cup \dots \cup E_k) \cap E_{k+1}). \quad (1)$$

Result follows by substituting required form for $P(E_1 \cup \dots \cup E_k)$, and re-writing

$$P((E_1 \cup \dots \cup E_k) \cap E_{k+1}) = P((E_1 \cap E_{k+1}) \cup \dots \cup (E_k \cap E_{k+1}))$$

(using distributivity) which is the union of k events, and hence can use the inductive hypothesis to re-write this final expression in the required form. Specifically, let $F_i = E_i \cap E_{k+1}$; then we have

$$\begin{aligned} P(F_1 \cup F_2 \cup \dots \cup F_k) &= \sum_i P(F_i) - \sum_i \sum_j P(F_i \cap F_j) + \dots (-1)^{k-1} P(F_1 \cap F_2 \cap \dots \cap F_k) \\ &= \sum_i P(E_i \cap E_{k+1}) - \sum_i \sum_j P(E_i \cap E_j \cap E_{k+1}) + \dots \\ &\quad (-1)^{k-1} P(E_1 \cap E_2 \cap \dots \cap E_k \cap E_{k+1}) \end{aligned}$$

as, for example,

$$F_i \cap F_j = (E_i \cap E_{k+1}) \cap (E_j \cap E_{k+1}) = E_i \cap E_j \cap E_{k+1}$$

Hence, using the inductive hypothesis to express

$$P(E_1 \cup \dots \cup E_k)$$

in (1) and adding in the final term, we complete the proof.

TUTORIAL SHEET WEEK 3 : SOLUTIONS

1. Try to verify Axioms (I), (II) and (III) within each framework.

RELATIVE FREQUENCY: $P(E) = \frac{n_E}{n}$ as $n \rightarrow \infty$.

(I) OK, as n_E/n always lies in $[0, 1]$.

(II) OK, as $n_\Omega = n$, so $P(\Omega) = n_\Omega/n = 1$.

(III) OK, as $n_{E \cup F} = (n_E + n_F)$ if E and F are disjoint, so $P(E \cup F) = (n_E + n_F)/n = n_E/n + n_F/n = P(E) + P(F)$ as $n \rightarrow \infty$.

CLASSICAL: Identical to RF case, but for finite n .

SUBJECTIVE:

(I) OK;

$$\begin{aligned} P(E) > 1 &\implies \text{pay more than } M \text{ to win } M ! \\ P(E) < 0 &\implies \text{paid to play !} \end{aligned}$$

(II) OK; if $E = \Omega$, must pay M so that neither side will be bound to lose.

(III) OK; Consider two bets on events E and F ,

$$\left. \begin{aligned} P(E) = p_E &\implies \text{pay } p_E M \text{ to win } M \text{ on } E \\ P(F) = p_F &\implies \text{pay } p_F M \text{ to win } M \text{ on } F \end{aligned} \right\} \implies \text{pay } (p_E + p_F) M \text{ to win } M \text{ on } E \text{ or } F$$

if E and F are mutually exclusive, so $P(E \cup F) = (p_E + p_F) = P(E) + P(F)$.

2. Free choice, but suggest

(a) $\Omega = \{\text{JAN, FEB, ..., DEC}\}$; equally likely outcomes ?

(b) $\Omega = \{x : 1889 \leq x \leq 2000\}$; relative frequency, subjective ?

(c) $\Omega = \{1001, 1002, \dots, 5999, 6000\}$; equally likely outcomes calculations give probabilities, but these outcomes are not equally likely.

(d) $\Omega = \{x : x \geq 0\}$; relative frequency ?

(e) $\Omega = \{x : -50 \leq x \leq 35\}$; relative frequency, subjective ?