

M1S : ASSESSED COURSEWORK 3 : SOLUTIONS

1. (i) To deduce \mathbb{X} , note first that we require $f_X(x)$ to be non-negative. The function $-\ln x$ is only defined when $x > 0$, and is only non-negative when $x \leq 1$. Hence $\mathbb{X} \equiv (0, 1]$ is largest potential range. Now we also require that

$$\int_{\mathbb{X}} f_X(x) dx = 1$$

and here

$$\begin{aligned} \int_{\mathbb{X}} f_X(x) dx &= \int_0^1 f_X(x) dx = \int_0^1 -\ln x dx = \int_0^1 -1 \cdot \ln x dx \\ &= [-x \ln x]_0^1 + \int_0^1 x \frac{1}{x} dx && \text{(by parts)} \\ &= 0 + 1 = 1 \end{aligned}$$

so the function integrates to 1 on $\mathbb{X} \equiv (0, 1]$. Hence we deduce that $(0, 1]$ is the appropriate range for the random variable, and we have by the identical argument that

$$F_X(x) = \int_0^x f_X(t) dt = \int_0^x -\ln t dt = x(1 - \ln x) \quad 0 < x \leq 1$$

[5 MARKS]

(ii) We have that

$$h_X(x) = \frac{f_X(x)}{1 - F_X(x)} \quad x > 0$$

which we may re-express, setting $y = F_X(x)$

$$h_X(x) = \frac{1}{1 - y} \frac{dy}{dx} \tag{1}$$

as, from the definition given we know that

$$\frac{dy}{dx} = \frac{d}{dx} F_X(x) = f_X(x)$$

Integrating both sides of (1) we have

$$H_X(x) = \int \frac{1}{1 - y} \frac{dy}{dx} dx = -\log(1 - y)$$

where $H_X(x)$ is the integral of function $h_X(x)$. On rearrangement,

$$y = F_X(x) = 1 - \exp\{-H_X(x)\}$$

from which we deduce finally that

$$F_X(x) = 1 - \exp\left\{-\int_0^x h_X(t) dt\right\}$$

More formally,

$$h_X(x) = \frac{f_X(x)}{1 - F_X(x)} \implies H_X(x) = \int_0^x h_X(t) dt = \int_0^x \frac{f_X(t)}{1 - F_X(t)} dt = [-\log(1 - F_X(t))]_0^x = -\log(1 - F_X(x))$$

and the results follows on rearrangement.

[5 MARKS]

2. We have for the cdf of X , for time $x > 0$

$$\begin{aligned}
 F_X(x) &= \mathbf{P}[X \leq x] = 1 - \mathbf{P}[X > x] \\
 &= 1 - \mathbf{P}[n\text{th event occurs } \textit{later} \text{ than time } x] \\
 &= 1 - \mathbf{P}[\textit{fewer than } n \text{ events occur in } [0, x]] \\
 &= 1 - \sum_{i=0}^{n-1} \frac{e^{-\lambda x} (\lambda x)^i}{i!}
 \end{aligned}$$

using the Poisson process model assumptions. On differentiation to get $f_X(x)$, we may differentiate term by term in the summation

$$\begin{aligned}
 f_X(x) &= - \sum_{i=0}^{n-1} \frac{\lambda^i}{i!} \frac{d}{dx} \{x^i e^{-\lambda x}\} \\
 &= - \sum_{i=0}^{n-1} \frac{\lambda^i}{i!} \{i x^{i-1} e^{-\lambda x} - \lambda x^i e^{-\lambda x}\} \\
 &= -e^{-\lambda x} \left[\lambda \sum_{i=1}^{n-1} \frac{(\lambda x)^{i-1}}{(i-1)!} - \lambda \sum_{i=0}^{n-1} \frac{(\lambda x)^i}{i!} \right]
 \end{aligned}$$

splitting the sum in two parts, noting that the first term in the first sum ($i = 0$) does not contribute to the sum, and then combining terms. Now note that

$$\sum_{i=1}^{n-1} \frac{(\lambda x)^{i-1}}{(i-1)!} = \sum_{j=0}^{n-2} \frac{(\lambda x)^j}{j!}$$

where $j = i - 1$, and so

$$\begin{aligned}
 f_X(x) &= -\lambda e^{-\lambda x} \left[\sum_{j=0}^{n-2} \frac{(\lambda x)^j}{j!} - \sum_{i=0}^{n-1} \frac{(\lambda x)^i}{i!} \right] \\
 &= \lambda e^{-\lambda x} \frac{(\lambda x)^{n-1}}{(n-1)!}
 \end{aligned}$$

as other terms cancel. Hence

$$f_X(x) = \frac{\lambda^n}{(n-1)!} x^{n-1} e^{-\lambda x} \quad x > 0$$

as required.

[10 MARKS]