

## M1S : EXERCISES 6

### PROPERTIES OF DISCRETE PROBABILITY DISTRIBUTIONS

1. Show that the function

$$\frac{1}{1+\lambda} \left( \frac{\lambda}{1+\lambda} \right)^x$$

for parameter  $\lambda > 0$  is a valid probability mass function for a discrete random variable  $X$  taking values on  $\{0, 1, 2, \dots\}$ , and find the corresponding cumulative distribution function.

2. Suppose that discrete random variable  $X$  has a *negative binomial* distribution, that is, the probability mass function of  $X$  is given by

$$f_X(x) = \binom{x-1}{n-1} \theta^n (1-\theta)^{x-n} \quad x = n, n+1, \dots$$

and zero otherwise, for  $n \geq 1$  and  $0 < \theta \leq 1$ . Recall that  $X$  corresponds to the number of binary trials required to obtain  $n$  successes.

Find the probability mass function of discrete random variable  $Y$  defined by  $Y = X - n$ .

3. It can be shown that, if  $X$  has a *Hypergeometric distribution*,  $X \sim \text{Hypergeometric}(N, R, n)$  (so that  $X$  is the count of the number of Type I items in a sample of size  $n$ ), then if  $N$  and  $R$  are large, so that in particular  $N, R \gg n$ , then

$$f_X(x) \approx \binom{n}{x} \theta^x (1-\theta)^{n-x} \quad x = 0, 1, \dots, n$$

and zero otherwise, where  $\theta = R/N$ , then  $X$  has an approximate Binomial distribution, with parameters  $n$  and  $\theta$ . Explain this result by considering sampling without replacement from a finite but large population.

4. A fair coin is tossed  $n$  times. Let  $X$  be the discrete random variable corresponding to the difference between the number of heads and the number of tails observed. Find the range and probability mass function of  $X$ .

5. If  $X$  has a Geometric distribution with parameter  $\theta$ , so that

$$f_X(x) = (1-\theta)^{x-1} \theta \quad x = 1, 2, 3, \dots$$

and zero otherwise, show that, for  $n, k \geq 1$ ,

$$P[ X = n + k \mid X > n ] = P[ X = k ]$$

This result is known as the *Lack of Memory* property.

6. For which values of  $k$  and  $\alpha$  are the following functions valid probability mass functions on the ranges given;

$$(i) f(x) = \frac{k}{x(x+1)}, \quad x = 1, 2, 3, \dots \quad (ii) f(x) = kx^\alpha, \quad x = 1, 2, 3, \dots$$

7. Consider a sequence of *Bernoulli* random variables  $X_1, \dots, X_n$  each with parameter  $\theta$  resulting from independent binary trials, so that

$$P[X = 0] = 1 - \theta \quad P[X = 1] = \theta$$

Find the probability distributions of the random variables

$$(i) Y = \text{Min} \{X_1, \dots, X_n\} \quad (ii) Z = \text{Max} \{X_1, \dots, X_n\}$$

[Hint: find the ranges of  $Y$  and  $Z$ , and consider  $P[ Y = 1 ]$ ,  $P[ Z = 0 ]$ ]

8. Consider an event  $E$  in sample space  $\Omega$ , and consider the *indicator function*,  $I_E$ , defined for  $\omega \in \Omega$  by

$$I_E(\omega) = \begin{cases} 1 & \omega \in E \\ 0 & \omega \notin E \end{cases}$$

Show that  $I_E$  defines a *Bernoulli* random variable.

Show furthermore that **any** discrete random variable can be expressed as a linear combination of indicator random variables.

9. The *Poisson distribution* is a probability model that is commonly used to specify probabilities for the outcomes of counting experiments. If  $X$  is a discrete random variable taking values on range  $\mathbb{X} = \{0, 1, 2, \dots\}$  with probability mass function  $f_X$  given by

$$f_X(x) = \mathbb{P}[X = x] = \frac{e^{-\lambda} \lambda^x}{x!} \quad x = 0, 1, 2, \dots$$

for parameter  $\lambda > 0$ , then  $X$  has a Poisson distribution with parameter  $\lambda$ , or  $X \sim \text{Poisson}(\lambda)$ .

(i) Verify that the Poisson distribution is a valid discrete probability distribution.

(ii) Suppose that  $Z$  is a discrete random variable with  $\mathbb{P}[Z = 0] = 0$ , but otherwise with  $\mathbb{P}[Z = z] \propto \mathbb{P}[X = z]$  for  $z \in \mathbb{R}$ . Find the probability mass function of  $Z$ .

(iii) Suppose now that  $X_1 \sim \text{Poisson}(\lambda_1)$  and  $X_2 \sim \text{Poisson}(\lambda_2)$  are **independent** random variables, so that

$$\mathbb{P}[(X_1 = x_1) \cap (X_2 = x_2)] = \mathbb{P}[X_1 = x_1] \mathbb{P}[X_2 = x_2] \quad \text{for all pairs } (x_1, x_2)$$

If random variable  $Y$  is defined by  $Y = X_1 + X_2$ , prove that  $Y \sim \text{Poisson}(\lambda_1 + \lambda_2)$ .

Hint: consider the partition

$$(Y = y) = \bigcup_{x_1=0}^{\infty} ((X_1 = x_1) \cap (X_2 = y - x_1)) \quad \text{for } y = 0, 1, 2, \dots$$

10. For a discrete random variable  $X$  with range  $\mathbb{X}$  and probability mass function  $f_X$ , define the **probability generating function** (or pgf) of  $X$ ,  $G_X$ , as a power series in  $t$  by

$$G_X(t) = \sum_{x \in \mathbb{X}} f_X(x) t^x$$

for values of  $t$  where the sum is convergent. The pgf is a function of  $t$  that is specific to each mass function  $f_X$ ; each  $f_X$  has a unique corresponding  $G_X$ .

(i) Find the pgf of  $X$  if  $X \sim \text{Poisson}(\lambda)$ .

(ii) Prove that if  $X_1$  and  $X_2$  are independent random variables, and  $Y = X_1 + X_2$ , then

$$G_Y(t) = G_{X_1}(t) G_{X_2}(t)$$